

Week 2 - L04

# Stochastic Gradient Descent (Examples)

CS 295 Optimization for Machine Learning

Ioannis Panageas

# Optimization in ML, SGD to the rescue

**Definition (Risk Minimization).** Let  $\ell(x, z) : \mathcal{X} \times \mathcal{Z} \rightarrow \mathbb{R}$  be a risk function and  $D$  some unknown distribution we can get samples from. We are interested in solving:

$$\min_{x \in \mathcal{X}} L(x), \text{ where } L(x) := \mathbb{E}_{z \sim D}[\ell(x, z)].$$

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Approach one:

1. Take **enough** (say  $n$ ) samples  $z_i$  **independently** and consider the **estimate**  $\bar{L}(x) := \frac{1}{n} \sum_i \ell(x, z_i)$ . By **Law of Large Numbers** this is a close enough with high probability.
2. Run a first order optimization algorithm (say GD) on  $\bar{L}(x)$ .

Remark:

If we do not know the form of  $\ell(x, z)$  and we only have oracle access it is not possible. Also many calculations per iteration...

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Approach one:

Or use SGD!

2. Run a first order optimization algorithm (say GD) on  $\bar{L}(x)$ .

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Approach two (SGD):

1. For each iteration  $t + 1$ , take a fresh sample  $z_t$  independently from  $z_1, \dots, z_{t-1}$  and consider the unbiased estimate  $\nabla_x \ell(x_t, z_t)$ .
2. Update  $x_{t+1} = x_t - \alpha_t \nabla_x \ell(x_t, z_t)$ .

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For  $t:=1$  to  $T$  do

1. sample  $z \sim D$ .
2. Pick  $v_t \in \partial \ell(x_t, z)$ .
3.  $x_{t+1} = x_t - \alpha_t v_t$ .

Return  $\frac{1}{T} \sum x_t$ .

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Remark:

What if  $L(x) = \frac{1}{n} \sum_{i=1}^n g_i(x)$ ?

Less Cost per iteration (but you need to “pay” the variance):

1. For each iteration  $t + 1$ , take uniformly at random **independently index  $i$**  from  $1, \dots, n$  and consider the **(unbiased estimate)**  $\nabla_x g_i(x)$ .
2. Update  $x_{t+1} = x_t - \alpha_t \nabla_x g_i(x)$ .

# An example (SGD approach)

**Definition (MLE for Gaussian).** Let  $z \sim \mathcal{N}(\mu, 1)$  and  $\ell(x, z) := -\log p_x(z)$  denotes the log-likelihood of  $\mathcal{N}(x, 1)$ . We *do not know*  $\mu$ . We are interested in solving:

$$\min_{x \in \mathbb{R}} \mathbb{E}_{z \sim \mathcal{N}(\mu, 1)} [-\log p_x(z)].$$

Any guesses what is the minimizer of the above?



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Of course  $x^* = \mu$ . Remarks on Maximum (log)-Likelihood:

1. Standard approach for parameter estimation of parametric families of distributions, i.e., create an **optimization** problem!
2. Under **assumptions**, Maximum (log) Likelihood Estimator is **consistent!**
3. Above boils down to  $\min_{x \in \mathbb{R}} \mathbb{E}_{z \sim \mathcal{N}(\mu, 1)} \left[ \frac{(z - x)^2}{2} \right]$ .
4. Let's do SGD...

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$$\min_{x \in \mathbb{R}} \mathbb{E}_{z \sim \mathcal{N}(\mu, 1)} \left[ \frac{(z - x)^2}{2} \right].$$

- The derivative is just  $(x - z)$  and  $\mathbb{E}[(x - z)^2] = 1 + (x - \mu)^2$ .
- The second derivative is 1, hence **1-strongly convex**.
- Start from  $x_0 = 0$ .
- At iteration  $t+1$ , get a **fresh** sample  $z_t$  and we have  $x_{t+1} = x_t - \alpha_t(x_t - z_t)$ .

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Recall for  $T = \Theta\left(\frac{1}{\epsilon} \log \frac{1}{\epsilon}\right)$  we get error  $\epsilon$ !

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- The derivative is just  $(x - z)$  and  $\mathbb{E}[(x - z)^2] = 1 + (x - \mu)^2$ .
- You can get  $\epsilon$ -close to  $\mu$  after  $\frac{1}{\epsilon^2} \ln \frac{1}{\epsilon^2}$  iterations! Not tight, why?
- Start from  $x_0 = 0$ .
- At iteration  $t+1$ , get a **fresh** sample  $z_t$  and we have  $x_{t+1} = x_t - \alpha_t(x_t - z_t)$ .

It is the empirical mean, i.e.,  $x_T = \frac{1}{T} \sum_i z_i$ !

# An example (SGD approach)

**Problem (Bias of a coin).** Assume you are given a coin that gives H with probability  $p \in (0, 1)$  and T with probability  $1 - p$ . *How many tosses* do you need to get an estimate  $\tilde{p}$  about  $p$  and be sure with probability 99% that  $|p - \tilde{p}| \leq \epsilon$ ?

Hint: Density  $f_p(z) = p^z (1 - p)^{1-z}$

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- A discrete probabilist will use Chernoff bounds or Chebyshev!
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We would like to solve (of course  $x^* = p$  is the solution but we don't know  $p$ )

$$\min_x \mathbb{E}[-z \log x - (1 - z) \log(1 - x)].$$

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Hint: Density  $f_p(z) = p^z(1 - p)^{1-z}$

- The derivative of  $\ell$  is just  $-\frac{z}{x} + \frac{(1-z)}{1-x} = \frac{x-z}{x(1-x)}$ , which is in absolute value at most  $\frac{1}{\epsilon}$  for  $x \in (\epsilon, 1 - \epsilon)$ .
- The second derivative of  $L$  is  $\frac{p}{x^2} + \frac{1-p}{(1-x)^2}$ , hence  $4(p - p^2)$ -strongly convex in  $(0, 1)$ .
- Start from  $x_0 = 1/2$ .
- At iteration  $t+1$ , get a **fresh** sample  $z_t$  and we have  $x_{t+1} = x_t - \alpha_t \frac{(x_t - z_t)}{x_t(1-x_t)}$ .

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Hint: Density  $f_p(z) = p^z(1 - p)^{1-z}$

- The derivative of  $\ell$  is just  $-\frac{z}{x} + \frac{(1-z)}{(1-x)} = \frac{x-z}{x(1-x)}$ , which is in absolute value  $\frac{1}{4(p-p^2)\epsilon^2}$ . You can get  $\epsilon$ -close to  $p$  after  $\frac{1}{4(p-p^2)\epsilon^2} \ln \frac{1}{\epsilon^2}$  iterations! Not tight, why?
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# A strange example.

**Problem (Mixture of Gaussians).** Assume you have access to i.i.d samples from  $z \sim \mathcal{N}(\mu, 1)$ . However, there is an adversary that with probability  $1/2$  corrupts  $z$  and gives you  $-z$ . Can you infer/estimate  $\mu$ ?

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$$\text{Need to solve: } \min_{x \in \mathbb{R}} \mathbb{E}_{z \sim \mathcal{N}(\mu, 1)} \left[ -\log \left( \frac{1}{2\sqrt{2\pi}} e^{(z-x)^2/2} + \frac{1}{2\sqrt{2\pi}} e^{(z+x)^2/2} \right) \right].$$

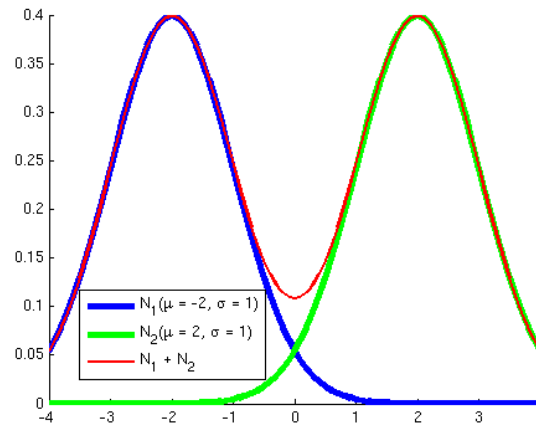
- Is it convex? **Exercise 5.**

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# Conclusion

- Examples on SGD:
  - MLE, testing bias of coin.
- Non-convex examples: Mixture of Gaussians
- Next week we will talk about **online learning/optimization!**